WWTP diagnosis based on robust principal component analysis

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Robust PCA





Application to hydraulic part of a wastewater treatment plant

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PCA

Maximization of the variance projections T = XP

- $T \in \Re^{N \times n}$: principal component matrix
- $P \in \Re^{n \times n}$: projection matrix

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Decomposition in eigenvalues/eigenvectors of the covariance matrix

$$\Sigma = \frac{1}{N-1} X^T X = P \Lambda P^T \quad \text{with} \quad P P^T = P^T P = I_n$$
$$\Sigma = \begin{bmatrix} P_\ell & P_{n-\ell} \end{bmatrix} \begin{bmatrix} \Lambda_\ell & 0\\ 0 & \Lambda_{n-\ell} \end{bmatrix} \begin{bmatrix} P_\ell^T\\ P_{n-\ell}^T \end{bmatrix}$$



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Principal part	×2
$\hat{X} = X C_{\ell}$ with $C_{\ell} = P_{\ell} P_{\ell}^{T}$	
Residual part	
$E = X - \hat{X}$ = X (I _n - C _ℓ) = X P _{n-ℓ} P ^T _{n-ℓ}	

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Determination of the number of principal components ℓ

Tharrault, Mourot, Ragot, Harkat (CRAN)

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Sensitive to outliers

Outliers: Data different from the normal operating conditions (faulty data, data obtained during shutdown or stratup periods or data issued form different operating mode)



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Robust PCA with respect to outliers

 \rightarrow Outliers detection and isolation



- n = 2 variables (x_1, x_2)
- $X = [x_1 x_2]$
- l = 1





- n = 2 variables (x_1, x_2)
- $X = [x_1 \ x_2]$
- l = 1
- Outliers 1 : green observation
- Outliers 2 : red observation





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Residual r(k)

 $r(k) = ||P_{n-\ell}^T(x(k) - \mu)||^2$

with
$$x(k)$$
 an observation
 μ the mean of the data X
 $P_{n-\ell}$ is the eigenvector matrix of the robust covariance matrix
corresponding to its $n-\ell$ smallest eigenvalues

The general scale-M estimator minimizes the following objective function with the constraint $P_{n-\ell}^{T}P_{n-\ell} = I_{n-\ell}$:

$$\frac{1}{N}\sum_{k=1}^{N}\rho\left(\frac{r(k)}{\hat{\sigma}}\right)$$

with the function ρ as the objective function, $\hat{\sigma}$ the robust scale of the residual $r_{(k)}$ calculated by minimising the following criterion:

$$\frac{1}{N}\sum_{k=1}^{N}\rho\left(\frac{r(k)}{\hat{\sigma}}\right) = \delta$$





Initialization with a robust covariance matrix

Robust covariance matrix

$$V = \frac{\sum_{i=1}^{N-1} \sum_{j=i+1}^{N} w(i,j) (x(i) - x(j)) (x(i) - x(j))^{T}}{\sum_{i=1}^{N-1} \sum_{j=i+1}^{N} w(i,j)}$$

where the weights w(i,j) themselves are defined by:

$$w(i,j) = \exp\left(-\frac{\beta}{2}(x(i)-x(j))^{T}\Sigma^{-1}(x(i)-x(j))\right)$$

with $\boldsymbol{\beta}$ turning parameter

For $\beta = 0$, $V = 2\Sigma$



Only robust to outliers with a projection onto the residual space

sacle M-estimator

The scale M-estimator maximizes the following criterion with the constraint $P_{\ell}^{T}P_{\ell} = I_{\ell}$:

$$\frac{1}{N}\sum_{k=1}^{N} \rho\left(\frac{||P_{\ell}^{T}(x(k) - \mu)||^{2}}{\hat{\sigma}}\right)$$

with $\hat{\sigma}$ the robust scale of the residual *r* and ρ the objective function.

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The reconstruction \hat{x}_R

Minimizing the influence of fault

$$\hat{x}_R(k) = x(k) - \Xi_R f_R$$

with x(k) : an observation

 f_R : the fault magnitude (unknown)

 Ξ_R : matrix of reconstruction directions

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For example, to reconstruct 2 variables (R = 2, 4) among 5 variables

$$\Xi_{R} = \left[\begin{array}{cccc} 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \end{array} \right]^{T}$$

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Estimation of the fault magnitude \hat{f}_R :

$$\hat{f}_R = \arg\min_{f_R} \{D_R(k)\}$$

with $D_R(k) = \hat{x}_R(k)^T P \Lambda^{-1} P^T \hat{x}_R(k)$



$$\hat{x}_{R}(k) = \left(I - \Xi_{R}(\Xi_{R}^{T}P\Lambda^{-1}P^{T}\Xi_{R})^{-1}\Xi_{R}^{T}P\Lambda^{-1}P^{T}\right) x(k)$$

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Condition of reconstruction:

Existence of $(\Xi_R^T P \Lambda^{-1} P^T \Xi_R)^{-1} \Rightarrow \text{matrix } \Xi_R^T P \Lambda^{-1} P^T \Xi_R$ of full rank



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To reconstruct a fault, it must be at least projected onto the principal space ($r \le \ell$) or onto the principal space ($r \le n - \ell$). This implies that the number of reconstructed variables *r* must respect the following inequality:

$$r \leq \max(n-\ell,\ell)$$

with r : Number of reconstructed variables

- ℓ : Number of principal components
- n: Number of variables



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The number of maximum reconstruction:

$$\sum_{r=1}^{\max(n-\ell,\ell)-1} \mathbb{C}'_r$$

with \mathbb{C}_n^r denotes the combination of *r* from *n*.



Fault detection indicator D_R

$$D_R(k) = \hat{x}_R(k)^T P \Lambda^{-1} P^T \hat{x}_R(k)$$

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Fault detection indicator D_R

$$D_R(k) = \hat{x}_R(k)^T P \Lambda^{-1} P^T \hat{x}_R(k)$$

Fault detection

A fault is detected, if:

$$D_R(k) > \gamma_{\alpha}^2$$

with γ_{α}^2 the detection threshold of indicator D_R ,

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Fault detection indicator D_R

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Fault detection

A fault is detected, if:

$$D_R(k) > \gamma_{\alpha}^2$$

with γ_{α}^2 the detection threshold of indicator D_R ,

Fault isolation

For the faulty observations, the faulty variables \hat{R} are determined as follows:

$$\hat{R} = \mathop{\mathrm{arg}}_{R \in \mathfrak{I}} D_R(k) < \gamma_{lpha}^2$$

with γ_{α}^2 the detection threshold of indicator D_R and \Im all combinations of possible reconstruction directions.

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Reduction of the number of reconstruction

Determination of the colinear direction projection

a global indicator K is built:

$$K(R_1, R_2) = \max\{(d(R_1, R_2), \tilde{d}(R_1, R_2)\}$$

with R_1 and R_2 correspond to sets of variable reconstruction and $d(R_1, R_2)$ distance between two sub-spaces onto principal space and $\tilde{d}(R_1, R_2)$ distance between two sub-spaces onto residual space

$$\begin{split} d(R_1, R_2) &= || \hat{\Xi}_{R_1} (\hat{\Xi}_{R_1}^T \hat{\Xi}_{R_1})^{-1} \hat{\Xi}_{R_1}^T - \hat{\Xi}_{R_2} (\hat{\Xi}_{R_2}^T \hat{\Xi}_{R_2})^{-1} \hat{\Xi}_{R_2}^T ||_2 \\ \tilde{a}(R_1, R_2) &= || \tilde{\Xi}_{R_1} (\hat{\Xi}_{R_1}^T \hat{\Xi}_{R_1})^{-1} \tilde{\Xi}_{R_1}^T - \tilde{\Xi}_{R_2} (\tilde{\Xi}_{R_2}^T \tilde{\Xi}_{R_2})^{-1} \tilde{\Xi}_{R_2}^T ||_2 \end{split}$$

with $\hat{\Xi}_{R_1} = \Lambda_{\ell}^{-1/2} P_{\ell}^T \Xi_{R_1}$, $\tilde{\Xi}_{R_1} = \Lambda_{n-\ell}^{-1/2} P_{n-\ell}^T \Xi_{R_1}$.

For example, to reconstruct 2 variables ($R_1 = 2, 4$) among 5 variables

$$\Xi_{R_1} = \left[\begin{array}{rrrr} 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \end{array} \right]^T$$



Algorithm for the determination of the detectable and isolable faults

r = 1

- Calculate for all available directions (R₁ ∈ ℑ and R₂ ∈ ℑ) the indicator K(R₁, R₂). If K(R₁, R₂) is equal to zero:
 - only a set of variables potentially faulty may be determined, i.e. the faulty variables are associated to the indices R₁ or R₂ or R₁ and R₂. Thus, it is only required to determine one direction, for example R₁.

If $K(R_1, R_2)$ is closed to zero:

• magnitude of the fault has to be important to ensure fault isolation.

Else the fault are isolable

- r = r + 1
- While $r < \max(\ell, n \ell)$ do to the step 2

Application to hydraulic part of a wastewater treatment plant



Description of the station



Construction of the data matrix

Dynamic process : Temporal lag Non linear process : Transformed variables

$$\begin{aligned} \mathbf{x}(k) &= \begin{bmatrix} H1(k) & H2(k) & H3(k) & Q5(k) & H6(k) \\ tanh((Q5(k-1)-550)./150) & H1(k-1) & (1) \\ H6(k-1) & C4(k) \end{bmatrix}^{\mathsf{T}} \end{aligned}$$

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The data matrix X is constituted of N observations of the vector x(k).



Construction of the model

4 principal components are determined



Measure and estimation of H1



Analysis of the reconstruction directions

The maximum number of reconstructions is then equal to 255 ($max(n-\ell,\ell)-1=4$) Κ R_1 1 2 3 4 5 6 7 8 9 0 0.94 0.99 1.00 0.99 0.98 0.97 0.99 1.00 1 $R_{2} = \begin{bmatrix} 1 \\ 2 \\ 3 \\ 4 \\ 5 \\ 6 \\ 7 \\ 8 \end{bmatrix}$ 0 0.52 1.00 0.98 1.00 1.00 1.00 1.00 0 0.99 1.00 1.00 1.00 1.00 0.99 1.00 0.99 1.00 1.00 0.33 0 0.82 0.96 0.90 0.99 0.80 1.00 1.00 0 1.00 0.97 0 0.97 Indicator K for r = 1

The number of useful reconstruction can be reduced to 202

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Fault detection



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Fault isolation

Fault index	Reconstruction direction
	under the detection threshold
16, 17	<i>D</i> ₁
3, 10, 11, 14, 15, 19, 20, 21	D3
7, 12	D _{1,6}
6, 8, 13, 18, 22	D _{3,9}
1, 2, 5	D _{1,3,4}
4	D _{3,7,9}
9	D _{1,2,3,7}

Table: Summary of fault isolation

Faults 16 and 17 are under the threshold when the first variable (H1(k)) is reconstructed. \Rightarrow variable H1(k) is faulty

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Conclusion

- Robust PCA with respect to outliers
 - -> directly applicable on data containing potential faults
- use of the principle of reconstruction and projection of the reconstructed data together
 -> outliers detection and isolation
- Reduction of the computational load
 > determination of the dectable and isolable faults
- Application to hydraulic part of a wastewater treatment plant

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